# On a new numerical analysis for the symmetric shortest queue problem 

Tayeb Lardjane, Rabah Messaci


#### Abstract

We consider a network of two M/M/1 parallel queues having the same poisonnian arrival stream with rate $\lambda$. Upon his arrival to the system a customer heads to the shortest queue and stays until being served. If the two queues have the same length, an arriving customer chooses one of the two queues with the same probability. Each duration of service in the two queues is an exponential random variable with rate $\mu$ and no jockeying is permitted between the two queues. A new numerical method, based on linear programming and convex optimization, is performed for the computation of the steady state solution of the system.


Keywords-steady state solution, matrix formulation, convex set, shortest queue, linear programming.

## I. Introduction

IN this paper we consider two identical $\mathrm{M} / \mathrm{M} / 1$ queues in parallel with a JSQ (joining the shortest queue) policy. The customers arrive to the system in accordance with a Poisson process of rate $\lambda$. Upon his arrival to the system, the customer joins the shortest queue and stays in the queue until being served. If the two queues have the same length, the customer joins one of the two queues with the same probability. The duration of each service is an exponential random variable with rate $\mu$. No jockeying is permitted between the two queues. The goal of the present paper is the numerical computation of the steady state probability $p(i, j)$ where $i$ is the number of customers in queue 1 and $j$ the number of customers in queue 2.

The shortest queue problem was initially proposed by Haight ([?]). A successful model for the symmetric case was given when the jockeying between the two queues is allowed. Kingman ([7]) obtained some asymptotic results for the joint steady state distribution of the number of customers in the two queues. Flatto and Mckean ([3]) used generating function techniques to give some limiting properties for the steady state probability. A numerical approach, using matrix geometric ([9]) technique was developed by Gertsbakh ([4]). Cohen and Boxma ([2]) showed that the analysis of the shortest queue problem can be reduced to the solution of a boundary value problem
Halfin ([6]) performed a linear programming method for the problem by giving a lower and an upper bound for the probability distribution of the total number of customers in the system. Zhao and Grassmann ([13]) used the Flatto and McKean ([3]) results and analysis tools in order to develop a numerical solution. Adan, Wessels and Zijm ([1]) showed

[^0]that the steady state distribution of the queue length is a mixture of product form distributions. Wang and Locker ([11]) presented a model where the state space of the related Markov process was truncated into banded arrays, so they derive the probability of queue length and the customer sojourn time. Yao and Knessl ([12]) considered the JSQ problem with two $\mathrm{M} / \mathrm{M} / \infty$ queues; they perform an analytical and a numerical computation of the steady state solution . Recently Tarabia ([10]) presented a solution for the problem when jockeying is permitted and the capacity of each queue is finite. In this paper, we present an easy computation method adapted for this type of problems, based on a new formulation of the steady state balance equations. The probabilities $p(i, j)$ are linear expressions of the diagonal probabilities $p(i, i)$. An algorithm based on a linear program will permit us to get a convex set which contains the last probabilities. This new method was succesfully tested for two infinite server parallel queues in [8]. The software available for the simplex method allows us to make a computation with a high precision for a large range of values of the intensity traffic $\rho=\frac{\lambda}{\mu}$.

## II. STEADY STATE ANALYSIS OF THE SYSTEM

We consider the two dimensional stochastic process $\left(X_{t}, Y_{t}\right)_{t}$ where $X_{t}$ (resp. $Y_{t}$ ) is the number of customers in the queue 1 (resp. queue 2 ) at time t . The process $\left(X_{t}, Y_{t}\right)_{t}$ is a recurrent positive Markov process if and only if $\rho=\frac{\lambda}{\mu}<2$ (Kingman 1961). For $(i, j) \in E=\mathbb{N} \times \mathbb{N}$; let $p(i, j)=$ $\lim _{t \rightarrow \infty} P\left(X_{t}=i, Y_{t}=j\right)$ the steady state probability for the $\stackrel{t \rightarrow \infty}{\text { process. The corresponding generator matrice is denoted by }}$ $Q=\left(q\left(e, e^{\prime}\right)\right)_{\left(e, e^{\prime}\right) \in \mathbb{N}^{2} \times \mathbb{N}^{2}}$

Notation 1: For $k \geq 1$ we define the following $(k+1) \times 1$ vectors and $(k+1) \times(k+1)$ matrices:

$$
\begin{gathered}
X_{2 k}=\left(\begin{array}{c}
p(2 k, 0) \\
p(2 k-1,1) \\
\cdot \\
\cdot \\
p(k+1, k-1) \\
p(k, k)
\end{array}\right) X_{2 k+1}=\left(\begin{array}{c}
p(2 k+1,0) \\
p(2 k, 1) \\
\cdot \\
\cdot \\
p(k+2, k-1) \\
p(k+1, k)
\end{array}\right) \\
A_{2 k}=\left(\begin{array}{ccccccc}
1 & 1 & 0 & 0 . & 0 & 0 & 0 \\
0 & 1 & 1 & 0 & 0 & 0 & 0 \\
0 & 0 & 1 & 1 & 0 & 0 & 0 \\
0 & 0 & 0 & 1 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 1 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 1 \\
0 & 0 & 0 & 0 & 0 & 0 & 1
\end{array}\right)
\end{gathered}
$$

# International Journal of Engineering, Mathematical and Physical Sciences <br> ISSN: 2517-9934 <br> Vol:6, No:8, 2012 

$$
A_{2 k+1}=\left(\begin{array}{ccccccc}
1 & 1 & 0 & 0 & 0 & 0 & 0 \\
0 & 1 & 1 & 0 & 0 & 0 & 0 \\
0 & 0 & 1 & 1 & 0 & 0 & 0 \\
0 & 0 & 0 & 1 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 1 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 1 \\
0 & 0 & 0 & 0 & 0 & 0 & 2
\end{array}\right)
$$

Notation 2: The $(k+1) \times 1$ vectors $B_{2 k}$ and $B_{2 k+1}$ are also defined as follows:
For $k=1$

$$
B_{2}=\binom{\left((1+\rho) \frac{\rho}{2}-\frac{\rho}{2}\right) p(0,0)}{p(1,1)}
$$

For $k \geq 2$

$$
B_{2 k}=\left(\begin{array}{c}
(1+\rho) p(2 k-1,0) \\
(2+\rho) p(2 k-2,1)-\rho p(2 k-2,0) \\
(2+\rho) p(2 k-3,2)-\rho p(2 k-3,1) \\
\cdot \\
\cdot \\
(2+\rho) p(k+1, k-2)-\rho p(k+1, k-3) \\
(2+\rho) p(k, k-1)-\rho p(k, k-2)-\alpha_{k} \\
p(k, k)
\end{array}\right)
$$

where $\alpha_{k}=\frac{\rho}{2} p(k-1, k-1)$
For $k \geq 1$

$$
B_{2 k+1}=\left(\begin{array}{c}
(1+\rho) p(2 k, 0) \\
(2+\rho) p(2 k-1,1)-\rho p(2 k-1,0) \\
(2+\rho) p(2 k-2,2)-\rho p(2 k-2,1) \\
\cdot \\
\cdot \\
(2+\rho) p(k+1, k-1)-\rho p(k+1, k-2) \\
(2+\rho) p(k, k)-2 \rho p(k, k-1)
\end{array}\right)
$$

Theorem 3: The steady state system of balance equations can be written as follows:
$A_{2 k} X_{2 k}=B_{2 k} \quad$ and $\quad A_{2 k+1} X_{2 k+1}=B_{2 k+1} \quad$ for $\quad k \geq 1$
Proof: For all $n \geq 1$, as showed in the transition diagram in figure 1; the system of balance equations expresses the probabilities $\{p(i, j), i+j=n+1\}$ in terms of the probabilities $\{p(i, j), i+j=n\}$ and $\{p(i, j), \quad i+j=n-1\}$. So we distinguish two cases:

Case 1: $\quad n=2 k \quad k \geq 2$
Because of the symmetry, the system of balance equations is reduced to:

$$
(\lambda+\mu) p(2 k-1,0)=\mu p(2 k, 0)+\mu p(2 k-1,1)
$$

for $0<i<k-1$ :

$$
\begin{align*}
(\lambda+2 \mu) p(2 k-1-i, i) & =\lambda p(2 k-1-i, i-1)  \tag{4}\\
& +\mu p(2 k-i, i)+\mu p(2 k-1-i, i+1)
\end{align*}
$$

$$
\begin{equation*}
(\lambda+\mu) p(2 k, 0)=\mu p(2 k+1,0)+\mu p(2 k, 1) \tag{3}
\end{equation*}
$$

or:
Case 2: $\quad n=2 k+1 \quad k \geq 1$
Again, because of the symmetry, the system of balance equations can be written as:

$$
\begin{equation*}
(\lambda+2 \mu) p(k, k)=2 \lambda p(k, k-1)+2 \mu p(k+1, k) \tag{2}
\end{equation*}
$$

for $0<i<k$

$$
\begin{aligned}
& (\lambda+2 \mu) p(k+i, k-i)=\lambda p(k+i, k-i-1) \\
& \quad+\mu p(k+i+1, k-i)+\mu p(k+i, k-i+1)
\end{aligned}
$$

and

$$
2 p(k+1, k)=(2+\rho) p(k, k)-2 \rho p(k, k-1)
$$

for $0<i<k$

$$
\begin{aligned}
& p(k+i+1, k-i)+p(k+i, k-i+1) \\
& \quad=(2+\rho) p(k+i, k-i)-\rho p(k+i, k-i-1)
\end{aligned}
$$

and

$$
\begin{equation*}
p(2 k, 1)+p(2 k+1,0)=(1+\rho) p(2 k, 0) \tag{5}
\end{equation*}
$$

Which is the formulation: $A_{2 k+1} X_{2 k+1}=B_{2 k+1}$
Proposition 4: For $j=1,2, . ., k+1$ the $j^{t h}$ component of the vectors $X_{2 k}$ (resp. $X_{2 k+1}$ ) is of the form:

$$
\begin{array}{r}
\quad\left(\sum_{i=1}^{i=k} \alpha_{i, j}^{(2 k)} x_{i}+\alpha_{0, j}^{(2 k)}\right) x_{0} \\
\left(\operatorname{resp} .\left(\sum_{i=1}^{i=k} \alpha_{i, j}^{(2 k+1)} x_{i}+\alpha_{0, j}^{(2 k+1)}\right) x_{0}\right)
\end{array}
$$

where $\alpha_{i, j}^{(2 k)}, \quad \alpha_{i, j}^{(2 k+1)} \in \mathbb{R}, \quad x_{i} x_{0}=p(i, i) \quad$ for $i \geq 1$ and $p(0,0)=x_{0}>0$

Proof: We use a recurrence argument. For $k=1$, we get from the formulas of theorem 1 :

$$
\begin{aligned}
X_{2} & =\binom{p(2,0)}{p(1,1)}=\left(\begin{array}{cc}
1 & -1 \\
0 & 1
\end{array}\right)\binom{\left((1+\rho) \frac{\rho}{2}-\frac{\rho}{2}\right) x_{0}}{x_{1} x_{0}} \\
& =\binom{\frac{1}{2} x_{0}\left(\rho^{2}-2 x_{1}\right)}{x_{1} x_{0}} \\
X_{3} & =\binom{p(3,0)}{p(2,1)}=\left(\begin{array}{cc}
1 & -\frac{1}{2} \\
0 & \frac{1}{2}
\end{array}\right)\binom{\frac{1}{2} \rho^{2}(\rho+1) x_{0}}{x_{0}\left(2 x_{1}+\rho x_{1}-\rho^{2}\right)} \\
& =\binom{\frac{1}{2} x_{0}(\rho+2)\left(\rho^{2}-x_{1}\right)}{\left(-\frac{1}{2}\right)\left(\rho^{2}-\rho x_{1}-2 x_{1}\right) x_{0}}
\end{aligned}
$$

with $x_{1} x_{0}=p(1,1)$. We note that the expressions of $X_{2}$ and $X_{3}$ gives the coefficients $\alpha_{1, j}^{(2)}, \quad \alpha_{1, j}^{(3)}$ for $j=1,2$.
Now, we assume that $X_{2 k-1}$ satisfies the proposition, then because of the expressions giving the vectors $X_{2 k}$ and $X_{2 k+1}$ ( $B_{2 k}$ is given in terms of the $X_{2 k-2}$ and $X_{2 k-1}$ components, $B_{2 k+1}$ is given in terms of the $X_{2 k}$ and $X_{2 k-1}$ components), the needed epressions of $X_{2 k}$ and $X_{2 k+1}$ are obtained. So the coefficients $\alpha_{i, j}^{(n)}$ are computed by a recursion.

Remark 5: This last proposition shows that if $i+j=n$ then $p(i, j)$ is expressed in terms of $p(0,0), p(1,1), p(2,2), \ldots \ldots p\left(\left[\frac{n}{2}\right],\left[\frac{n}{2}\right]\right) \quad$ where $[x]$ is an integer having the property: $[x] \leq x<[x]+1$

Notation 6: In the following sections, for each integer $n$, we introduce the real numbers denoted $\left\{r_{n}(i, j), i+j \leq n\right\}$ satisfying the system of equilibrium equations

$$
\begin{equation*}
r_{n}(e) \sum_{e^{\prime} \neq e} q\left(e, e^{\prime}\right)=\sum_{e^{\prime} \neq e} r_{n}\left(e^{\prime}\right) q\left(e^{\prime}, e\right) \tag{6}
\end{equation*}
$$

where $(e)=(i, j) ; i+j \leq n-1$; so we put:
where $(e)=(i, j) ; i+j \leq n-1$, so we put:
$r_{n}(0,0)=y_{0}^{(n)}, r_{n}(0,1)=r_{n}(1,0)=\frac{\rho}{2} r_{n}(0,0), r_{n}(i, i)=$

$$
\begin{gathered}
y_{i}^{(n)} y_{0}^{(n)} \text { for } i \geq 1 \\
Y_{2 k}^{(n)}=\left(\begin{array}{c}
r_{n}(2 k, 0) \\
r_{n}(2 k-1,1) \\
\cdot \\
\cdot \\
r_{n}(k+1, k-1) \\
r_{n}(k, k)
\end{array}\right) Y_{2 k+1}^{(n)}=\left(\begin{array}{c}
r_{n}(2 k+1,0) \\
r_{n}(2 k, 1) \\
\cdot \\
\cdot \\
r_{n}(k+2, k-1) \\
r_{n}(k+1, k)
\end{array}\right) \\
D_{2}^{(n)}=\binom{\left((1+\rho) \frac{\rho}{2}-\frac{\rho}{2}\right) r_{n}(0,0)}{r_{n}(1,1)}
\end{gathered}
$$

For $k \geq 2$ and $2 k \leq n$

$$
=\left(\begin{array}{c}
(1+\rho) r_{n}(2 k-1,0) \\
(2+\rho) r_{n}(2 k-2,1)-\rho r_{n}(2 k-2,0) \\
(2+\rho) r_{n}(2 k-3,2)-\rho r_{n}(2 k-3,1) \\
\cdot \\
\cdot \\
(2+\rho) r_{n}(k+1, k-2)-\rho r_{n}(k+1, k-3) \\
(2+\rho) r_{n}(k, k-1)-\rho r_{n}(k, k-2)-\beta_{k} \\
r_{n}(k, k)
\end{array}\right)
$$

where $\beta_{k}=\frac{\rho}{2} r_{n}(k-1, k-1)$
For $k \geq 1$ and $2 k+1 \leq n$
For $k \geq 1$ and $2 k+1 \leq n$

$$
D_{2 k+1}^{(n)}=\left(\begin{array}{c}
(1+\rho) r_{n}(2 k, 0) \\
(2+\rho) r_{n}(2 k-1,1)-\rho r_{n}(2 k-1,0) \\
(2+\rho) r_{n}(2 k-2,2)-\rho r_{n}(2 k-2,1) \\
\cdot \\
\cdot \\
(2+\rho) r_{n}(k+1, k-1)-\rho r_{n}(k+1, k-2) \\
(2+\rho) r_{n}(k, k)-2 \rho r_{n}(k, k-1)
\end{array}\right)
$$

The vectors $Y_{2 k}^{(n)}$ and $Y_{2 k+1}^{(n)}$ are then defined by the recursive formulas:
$A_{2 k} Y_{2 k}^{(n)}=D_{2 k}^{(n)} \quad$ and $\quad A_{2 k+1} Y_{2 k+1}^{(n)}=D_{2 k+1}^{(n)} \quad$ for $\quad k \geq 1$ We also have: the $(k+1)$ components of the vectors $Y_{2 k}^{(n)}$ (resp $Y_{2 k+1}^{(n)}$ ) are of the form:

$$
\begin{aligned}
& \left(\sum_{i=1}^{i=k} \alpha_{i, j}^{(2 k)} y_{i}^{(n)}+\alpha_{0, j}^{(2 k)}\right) y_{0}^{(n)} \\
& \left(\operatorname{resp} \cdot\left(\sum_{i=1}^{i=k} \alpha_{i, j}^{(2 k+1)} y_{i}^{(n)}+\alpha_{0, j}^{(2 k+1)}\right) y_{0}^{(n)}\right)
\end{aligned}
$$

where $\alpha_{i, j}^{(2 k)}, \quad \alpha_{i, j}^{(2 k+1)} \in \mathbb{R}, \quad 1 \leq j \leq(k+1)$
Proposition 7: For $n \geq 2$, let the $\mathbb{R}^{\left[\frac{n}{2}\right]}$ sub set $C_{n}$ defined as follows:
$C_{n}=\left\{\begin{array}{c}\left(y_{1}, y_{2}, \ldots \ldots, y\left[\frac{n}{2}\right]\right) / y_{l}>0 \text { for } l=1,2, \ldots\left[\frac{n}{2}\right], \\ r(i, j)>0 \text { if } i+j=n\}\end{array}\right.$

$$
r(i, j)>0 \text { if } i+j=n\}
$$

then for $Y^{(n)}=\left(y_{1}^{(n)}, y_{2}^{(n)}, \ldots \ldots, y^{(n)}\left[\begin{array}{l}\left.\frac{n}{2}\right]\end{array}\right) \in C_{n}\right.$; the set
of real numbers $\left\{r_{n}(i, j) ; i+j \leq n\right\}$ (obtained in terms of components of the vector $Y^{(n)}$ in $C_{n}$ ) is a positive solution for the system of equilibrium equations (6)

Proof: $C_{n}$ is nonempty while it contains the vector $\left(p(1,1), p(2,2), \ldots \ldots p\left(\left[\frac{n}{2}\right],\left[\frac{n}{2}\right]\right)\right)$ up to a multiplicative factor and if $y_{i}^{(n)}=\frac{p(i, i)}{p(0,0)}$ for $1 \leq i \leq\left[\frac{n}{2}\right]$ and $y_{0}^{(n)}=p(0,0)$ the corresponding $r_{n}(i, j), i+j=n$ are exactly $p(i, j) ; i+j=n$.

The condition $r_{n}(i, j)>0$ if $i+j=n$ gives for those real numbers the properties of the steady state probabilities $p(i, j) ; i+j=n$. then $\left\{r_{n}(i, j) ; i+j \leq n-1\right\}$ ( from which we build up $\left\{r_{n}(i, j) ; i+j=n\right\}$ over the system of equilibrium equations (6) ; with $y_{i}^{(n)}$ playing the role of $\left.p(i, i), 1 \leq i \leq\left[\frac{n}{2}\right]\right)$ have the same properties than $\{p(i, j) ; i+j \leq n-1\}$. So if $y_{0}^{(n)}>0$ and $\left(y_{1}^{(n)}, y_{2}^{(n)}, \ldots \ldots, y^{(n)}\left[\frac{n}{2}\right]\right) \in C_{n}$ the builded system of real numbers $\left\{r_{n}(i, j) ; i+j \leq n\right\}$ is a positive solution for the equilibrium equations (6).

Corollary 8: If we note:
$S_{n}=\left\{\begin{array}{c}\left(y_{1}(n), y_{2}(n), \ldots \ldots, y\left[\frac{n}{2}\right](n), \ldots \ldots . .\right) \text { so that } y_{l}>0 \\ \forall l \geq 1, r(i, j)>0 \text { for } i+j=n\}\end{array}\right.$

$$
\forall l \geq 1, \quad r(i, j)>0 \text { for } i+j=n\}
$$

then $\left(S_{n}\right)_{n}$ is a decreasing sequence of sets $\left(S_{n} \subset S_{n-1}\right)$ and the limit $\cap S_{n}$ is so that $\lim _{n \rightarrow+\infty}\left(y_{i}^{(n)} y_{0}^{(n)}\right)_{i \geq 1}$ is exactly the entire diagonal probabilities $p(i, i) i \geq 1$.

Proof: Let $\left(y_{1}^{(n)}, y_{2}^{(n)}, \ldots \ldots ., y_{\left[\frac{n}{2}\right]}^{(n)}, \ldots \ldots ..\right) \in S_{n}$, we note first that the components $y_{l}^{(n)}$ for $l>\left[\frac{n}{2}\right]$ are free from the constraints $r_{n}(i, j)>0$ for $i+j=n$ then those components are identic for the two sets $S_{n}$ and $S_{n-1}$. While the components $\left(y_{1}^{(n)}, y_{2}^{(n)}, \ldots \ldots, y^{(n)}\left[\frac{n}{2}\right]\right)$ have also to fulfil the constraints $\left\{r_{n}(i, j)>0 ; i+j=n-1\right\}$ (previous proposition) then $S_{n} \subset S_{n-1}$. So, when $n$ goes to the infinity and due to the unicity of the positive solution for the infinite linear system of balance equations (the related Markov process is ergodic), the set $S_{n}$ has for limit the single point of $\mathbb{R}^{\mathbb{N}}$ which is the diagonal probabilities $(p(i, i))_{i \geq 1}$ up to a multiplicative factor.

Remark 9: All element of $C_{n}$ can be viewed as an approximation of the vector $(p(1,1), \ldots$.
..... $\left.p\left(\left[\frac{n}{2}\right],\left[\frac{n}{2}\right]\right)\right)$ up to a multiplicative factor.
Remark 10: While for all $n, p(i, j) ; i+j=n$ is expressed in terms of $p(0,0), p(1,1), p(2,2)$, $\ldots . ., p\left(\left[\frac{n}{2}\right],\left[\frac{n}{2}\right]\right)$, then the first step of computation is


Fig. 2. Evolution of the feasible regions
to get an approximation of those last probabilities by $y_{0}^{(n)}, y_{0}^{(n)} y_{1}^{(n)}, \ldots \ldots y_{0}^{(n)} y_{\left[\frac{n}{2}\right]}^{(n)}$ respectively, $y_{0}^{(n)}$ is a multiplicative factor and $y_{1}^{(n)}, \ldots \ldots y_{\left[\frac{n}{2}\right]}^{(n)}$ are computed under the constraints given in the definition of $C_{n}$. We then choose $n$ suficiently large in order to get a good approximation. In other words $y_{1}^{(n)} y_{0}^{(n)}, y_{2}^{(n)} y_{0}^{(n)}, \ldots \ldots \ldots ., y_{n}^{(n)} y_{0}^{(n)}$ are almost the diagonal probabilities: $p(1,1), \ldots, p(n, n)$. In order to show the geometric behavior of the set $C_{n}$ in the plane we choose $\rho=1.5$ and we sketch $C_{4}, C_{5}$ and the projection of $C_{6}$ on the plane containing the first two coordinates $\left(y_{1}, y_{2}\right)$ (Figure 2)

## III. Computation methodology

In practice we need a finite number of the the probabilities $\{p(i, i)\}$.The sequence $\{p(i, i)\}_{i}$ is decreasing after some integer $K$ (see [13]) and have the limit 0 (see Figure 3). The problem is then reduced to evaluate an integer $N$ large enough so that the sum $\sum_{\{(i, j) / i+j \leq N\}} p(i, j)$ is very close to 1. To do this, we use the bounds for $\sum_{n=N}^{\infty} \pi_{n}$ where $\pi_{n}=\sum_{\{(i ; j) / i+j=n\}} p(i, j)$ given in Halfin ([6]). We then have For $\sigma=\frac{\ln \frac{2+\rho}{(2-\rho)(1+\rho)}}{\ln (1+\rho)}$ and:

$$
\begin{array}{ll}
\sqrt{2} \leq \rho<2, & N \geq \sigma+1 \quad \sum_{n=N}^{\infty} \pi_{n} \leq\left(\frac{\rho}{2}\right)^{N-\sigma} \\
1 \leq \rho<\sqrt{2}, & N \geq 2 \quad \sum_{n=N}^{\infty} \pi_{n} \leq\left(\frac{\rho}{2}\right)^{N} \frac{2+\rho}{1+\rho}
\end{array}
$$

For $0<\rho<1$, we have:

$$
\begin{aligned}
\pi_{n} & \leq\left(\frac{\rho}{2}\right)^{n} \frac{1}{2(1+\rho)}\left(2+\rho-\frac{\rho^{2}(1+\rho)^{n}}{(1+\rho)^{n}-1}\right) \\
& +\frac{\rho^{2}}{2(1+\rho)} \frac{1}{(1+\rho)^{n}-1}
\end{aligned}
$$

which implies that:

$$
\begin{aligned}
\pi_{n} & \leq\left(\frac{\rho}{2}\right)^{n}+\frac{1}{(1+\rho)^{n}} \\
\text { so } \sum_{n=N}^{\infty} \pi_{n} & \leq\left(\frac{\rho}{2}\right)^{N} \frac{1}{1-\frac{\rho}{2}}+\left(\frac{1}{1+\rho}\right)^{N-1} \frac{1}{\rho}
\end{aligned}
$$

This last bound is not intersting if $\rho$ is close to 0 . In this case, we considere the intensity flows for a Markov process in equilibrium as invoked in Halfin [6] and having the form:

$$
\begin{aligned}
& \lambda \pi_{n-1}=\mu \pi_{n}^{-}+2 \mu \pi_{n}^{+} \text {where } \pi_{n}^{-}=2 p(n, 0) \text { and } \pi_{n}^{+}= \\
& \{i+j=n, i>0, j>0\} \\
& p(i, j) \quad\left(\pi_{n}=\pi_{n}^{-}+\pi_{n}^{+}\right), \text {we get: } \\
& \pi_{n}=\frac{\lambda}{\frac{\mu \pi_{n}^{-}+2 \mu \pi_{n}^{+}}{\pi_{n}}} \pi_{n-1}=\frac{\lambda}{\mu\left(1+\frac{\pi_{n}^{+}}{\pi_{n}}\right)} \pi_{n-1} \quad \text { so, } \pi_{n} \leq
\end{aligned}
$$

$\rho^{n}$ which is more intersting if $0<\rho<1$, with a special look for $\rho$ close to 0 . Thus: $\sum_{n=N}^{\infty} \pi_{n} \leq \frac{\rho^{N}}{1-\rho}$.

So, for a given precision $\epsilon$, the computation of the integer $N$ satisfying $\sum_{n=N+1}^{\infty} \pi_{n} \leq \varepsilon$ can be done easily. As an example for $\rho=1$ and $\epsilon=10^{-10}$ we find $N=34$. We then need $p(i, i) 1 \leq i \leq 17$. The determination of $N$ can also be
regarded as the stopping rule for the computation algorithm described below. We note that the upper bound $\epsilon$ of the error is made on the total sum of probabilities. We see further in the numerical results that $\epsilon$ is around the upper bound of the error made between the computed value of $p(i, i)$ and $p(i, i)$ itself.


Fig. 3. Evolution of the diagonal probabilities

## IV. COMPUTATION OF THE STEADY STATE PROBABILITIES

 AND NUMERICAL RESULTS
## A. Computation of $y_{i}^{(N)} 1 \leq i \leq\left[\frac{N}{2}\right]$

We recall that the integer $N$ is so that all the probabilities $\{p(i, j) / i+j \geq N+1\}$, having a total sum smaller than a given $\epsilon$, are neglected.
While the $\left[\frac{N}{2}\right]+1$ components of the vector $Y_{2\left[\frac{N}{2}\right]}^{(N)}$ are of the forme:

$$
\left(\sum_{i=1}^{i=\left[\frac{N}{2}\right]} \alpha_{i, j}^{(N)} y_{i}^{(N)}+\alpha_{0, j}^{(N)}\right) y_{0}^{(N)} ; 1 \leq j \leq\left[\frac{N}{2}\right]+1
$$

then the unknowns $\left\{y_{i}^{(N)} ; 1 \leq i \leq\left[\frac{N}{2}\right]\right\}$ are in the set solution of the system of linear inequalities:

$$
\begin{equation*}
\sum_{i=1}^{i=\left[\frac{N}{2}\right]} \alpha_{i, j}^{(N)} y_{i}^{(N)}+\alpha_{0, j}^{(N)}>0 \text { and } y_{i}^{(N)}>0 \tag{7}
\end{equation*}
$$

for $1 \leq j \leq\left[\frac{N}{2}\right]+1$ and $1 \leq i \leq\left[\frac{N}{2}\right]$.
We use the simplex algorithm to obtain a lower and an upper bound for each $y_{i}^{(N)}$ noted respectively $y_{i, \text { min }}^{(N)}$ and $y_{i, \text { max }}^{(N)}$. Because of the unicity due to the ergodicity of the related Markov process, those bounds are almost equal for a large value of $N$, as to be seen in further computation (see table 1). While $y_{i}^{(N)}>0$, we can take the objective function $\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}$ ( or any linear function of $\left\{y_{i}^{(N)} ; 1 \leq i \leq\left[\frac{N}{2}\right]\right\}$ ) and the constraints (7). While $\sum p(i, i)=\frac{1}{1+\rho}$ (see [6]), we then compute $y_{0}^{(N)}$ as a constant of normalisation by: $y_{0}^{(N)}\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right)=\frac{1}{1+\rho} \quad$ or $y_{0}^{(N)}=$ $\frac{1}{(1+\rho)\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right)}$ where $\left(y_{i}^{(N)}\right)_{1 \leq i \leq\left[\frac{N}{2}\right]}$ is any solu-
tion for (7). We can take the arithmetic mean: $y_{i}^{(N)}=$ $\frac{1}{2}\left(y_{i, \text { min }}^{(N)}+y_{i, \text { max }}^{(N)}\right)$

## B. Algorithm description

Step 1: set up a value of $\rho$.
Step 2: set up a precision $\epsilon$ and compute the corresponding value of $N$.
Step 3: put $y_{i}^{(N)} \cdot y_{0}^{(N)}=r_{N}(i, i) 1 \leq i \leq\left[\frac{N}{2}\right]$ and get (by using a formal calculus) the $\left[\frac{N}{2}\right]+1$ components of the vector $Y_{2\left[\frac{N}{2}\right]}^{(N)}$ denoted by $\sum_{i=1}^{i=\left[\frac{N}{2}\right]} \alpha_{i, j}^{\left(2\left[\frac{N}{2}\right]\right)} y_{i}^{(N)}+\alpha_{0}^{\left(2\left[\frac{N}{2}\right]\right)} 1 \leq$

ISSN: 2517-9934
Vol:6, No:8, 2012
$j \leq\left[\frac{N}{2}\right]+1$. This can be done whithout solving the formal systems of linear equations $A_{n} Y_{n}^{(N)}=D_{n}^{(N)}$ for $n=2 k$ and $n=2 k+1$. It's due to the fact that the inverse of the matrix:

$$
\left(\begin{array}{cccccc}
1 & 1 & 0 & . & . & 0 \\
0 & 1 & 1 & 0 & . & . \\
. & 0 & 1 & 1 & 0 & . \\
. & . & 0 & 1 & 1 & 0 \\
. & . & . & 0 & 1 & 1 \\
0 & . & . & . & 0 & 1
\end{array}\right)
$$

has one of the two forms:

$$
\begin{gathered}
\left(\begin{array}{ccccc}
1 & -1 & 1 & -1 & 1 \\
0 & 1 & -1 & 1 & -1 \\
\cdot & 0 & 1 & -1 & 1 \\
. & \cdot & 0 & 1 & -1 \\
0 & \cdot & \cdot & 0 & 1
\end{array}\right) \\
\text { or } \quad\left(\begin{array}{cccccc}
1 & -1 & 1 & -1 & 1 & -1 \\
0 & 1 & -1 & 1 & -1 & 1 \\
\cdot & 0 & 1 & -1 & 1 & -1 \\
. & \cdot & 0 & 1 & -1 & 1 \\
. & \cdot & \cdot & 0 & 1 & -1 \\
0 & \cdot & \cdot & \cdot & 0 & 1
\end{array}\right)
\end{gathered}
$$

and the inverse of the matrix:

$$
\left(\begin{array}{cccccc}
1 & 1 & 0 & . & . & 0 \\
0 & 1 & 1 & 0 & . & . \\
. & 0 & 1 & 1 & 0 & . \\
. & . & 0 & 1 & 1 & 0 \\
. & . & . & 0 & 1 & 1 \\
0 & . & . & . & 0 & 2
\end{array}\right)
$$

has one of the two forms:

$$
\begin{array}{r}
\left(\begin{array}{ccccc}
1 & -1 & 1 & -1 & \frac{1}{2} \\
0 & 1 & -1 & 1 & -\frac{1}{2} \\
. & 0 & 1 & -1 & \frac{1}{2} \\
. & \cdot & 0 & 1 & -\frac{1}{2} \\
0 & \cdot & . & 0 & \frac{1}{2}
\end{array}\right) \\
o r\left(\begin{array}{cccccc}
1 & -1 & 1 & -1 & 1 & -\frac{1}{2} \\
0 & 1 & -1 & 1 & -1 & \frac{1}{2} \\
. & 0 & 1 & -1 & 1 & -\frac{1}{2} \\
\cdot & \cdot & 0 & 1 & -1 & \frac{1}{2} \\
. & \cdot & . & 0 & 1 & -\frac{1}{2} \\
0 & \cdot & . & . & 0 & \frac{1}{2}
\end{array}\right)
\end{array}
$$

So in the computation program we put: $Y_{n}^{(N)}=A_{n}^{-1} D_{n}^{(N)}$ for $n=2 k$ or $2 k+1$.
Step 4: use the simplex method with objective function $i=\left[\frac{N}{2}\right]$
$\sum_{i=1}^{2} y_{i}^{(N)}$ and constraints:

$$
\sum_{i=1}^{i=\left[\frac{N}{2}\right]} \alpha_{i, j}^{(m)} y_{i}^{(N)}+\alpha_{0, j}^{(m)} \geq 0 \quad 1 \leq m \leq\left[\frac{N}{2}\right]+1
$$

$$
y_{i}^{(N)} \geq 0 \quad 1 \leq i \leq\left[\frac{N}{2}\right] .
$$

Step 5: get a lower and an upper bound for $y_{i}^{(N)}$ denoted respectively $y_{i, \min }^{(N)}$ and $y_{i, \max }^{(N)}$ and put $y_{i}^{(N)}=$ $\frac{1}{2}\left(y_{i, \max }^{(N)}+y_{i, \text { min }}^{(N)}\right)$. So the set solution is a convex set denoted $C_{N}^{\prime}$.
Step 6: return to the matrix formulation for the computation of $p(i, j) \quad i+j \leq N$.

## C. Numerical results and error analysis

Proposition 11: If we note $e_{0}$ and $e_{i}$ the error computation made on $p(0,0)$ and $\frac{p(i, i)}{p(0,0)}$ respectively then:

$$
e_{0} \leq \frac{\sum_{i=1}^{\left[\frac{N}{2}\right]}\left(y_{i, \max }^{(N)}-y_{i, \min }^{(N)}\right)}{(1+\rho)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, \min }^{(N)}\right)^{2}}+\frac{\epsilon\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, \max }^{(N)}\right)}{\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, \min }^{(N)}\right)^{2}}=
$$

$\epsilon_{0}$ and $\quad e_{i} \leq\left(y_{i, \text { max }}^{(N)}-y_{i, \text { min }}^{(N)}\right)=\epsilon_{1, i}$

is a sequence of real positive numbers having the property: $y_{i, \text { min }}^{(N)} \leq y_{i, 0}^{(N)} \leq y_{i, \max }^{(N)} 1 \leq$ $i \leq N$ for $\left(y_{1}^{(N)}, y_{2}^{(N)}, \ldots \ldots, y_{\left[\frac{n}{2}\right]}^{(N)}\right) \in C_{n}$. Then $p(0,0)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, 0}^{(N)}\right)=\frac{1}{1+\rho}-\alpha_{1}$ which leads to:

$$
p(0,0)=\frac{1-(1+\rho) \alpha_{1}}{(1+\rho)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, 0}^{(N)}\right)} \text { with } 0<\alpha_{1}<\epsilon \text { and }
$$

the integer $N$ satisfying: $\sum_{n=N+1}^{\infty} \pi_{n} \leq \epsilon$. So, the computed value of $p(0,0)$ is $y_{0}^{(N)}=\frac{1}{(1+\rho)\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right)}$ where
the sequence $\left(y_{i}^{(N)}\right)_{1 \leq i \leq\left[\frac{N}{2}\right]}$ is chosen in the set $C_{n}$.

We then have

$$
\begin{aligned}
& \left\lvert\, \begin{array}{l}
\left|p(0,0)-y_{0}\right|= \\
(1+\rho)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, 0}^{(N)}\right)
\end{array} \frac{\left.(1+\rho)\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right) \right\rvert\,}{\left(1-(1+\rho) \alpha_{1}\right)\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right)-\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, 0}^{(N)}\right)}\right. \\
& =\left|\frac{(1+\rho)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, 0}^{(N)}\right)\left(1+\sum_{1}^{\left[\frac{N}{2}\right]} y_{i}^{(N)}\right)}{\left(1+\sum_{i=1}^{[N)} y_{i, \max }^{(N)}\right)}\right| \\
& \leq \frac{\sum_{i=1}^{\left[\frac{N}{2}\right]}\left(y_{i, \max }^{(N)}-y_{i, \min }^{(N)}\right)}{(1+\rho)\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, \min }^{(N)}\right)^{2}}+\frac{1}{\left(1+\sum_{i=1}^{\left[\frac{N}{2}\right]} y_{i, \min }^{(N)}\right)^{2}}
\end{aligned}
$$

While the true value of $p(i, i)$ is given by: $p(i, i)=$ $p(0,0)\left(\lim _{N \rightarrow+\infty} y_{i}^{(N)}\right)$ where: $y_{i, \min }^{(N)} \leq y_{i}^{(N)} \leq y_{i, \max }^{(N)}$ for any $i$ and $N$ such that $1 \leq i \leq\left[\frac{N}{2}\right]$, then $y_{i, \min }^{(N)} \leq \frac{p(i, i)}{p(0,0)} \leq y_{i, \max }^{(N)}$
We denote further $\epsilon_{1}=\max \epsilon_{1, i}$ the upper bound of the error commited on the probabilities $p(i, i)$.

It is well known that the convergence to the solution is fast for small values of $\rho(\rho<1)$. More $\rho$ is close to 2 , more the convergence is slow. This can be seen in the plots made in figure 3. We give then a sample of results obtained for different values of $\rho$. After each computation of a system of probabilities $p(i, j)$ we add a table in which we indicate the values of:

- $\epsilon$ of the error made on the total sum of probabilities
- $\epsilon_{0}$ the upper bound of the error made on the computation of $p(0,0)$
- $\epsilon_{1}$ the upper bound of the error made on the computation of the $\frac{p(i, i)}{p(0,0)} i \geq 1$
- The integer $N$ having the property $\sum_{i+j \geq N+1} p(i, j) \leq \epsilon$
- The computed sum $\sum_{i+j \leq N} p(i, j)$

While the diagonal probabilities are of a special interest we also give them for $\rho=0.2$ and $\rho=1.9$. Some other cases of computation are made in further tables with values of $\rho$ close to 2 .

Results comments: The obtained results match those given in [13]. The powerfulness of the method described in the present paper allows us to make computations for all values of $\rho$ and this is done for $\rho=0.2$ and $\rho=1.9$. In addition to the precision with which are obtained the probabilities $p(i, j)$, this method allows us (by using the bounds given in [6]) to make an easy choice of the finite state space $E_{\epsilon}$ in which the computation is made. If $E$ is the whole state space of the process $\left(X_{t}, Y_{t}\right)_{t}$ then $E-E_{\epsilon}$ represents the missing states with $P\left(E-E_{\epsilon}\right) \leq \epsilon$.

ISSN: 2517-9934
Vol:6, No:8, 2012

TABLE I: SOME SIDES OF $\prod_{i=1}^{\left[\frac{N}{2}\right]}\left[y_{i, \text { min }}^{(N)}, y_{i, \text { max }}^{(N)}\right]$ FOR $\rho=1.5$

|  |  | $\left[\frac{N}{2}\right]=60$ |
| :--- | :--- | :--- |
| $i$ | $y_{i, \max }^{(N)}$ | $y_{i, \min }^{(N)}$ |
| 1 | 0.91576961624328916380 | 0.91576961624328916379 |
| 10 | 0.00581353037003270174 | 0.00581353037003270100 |
| 20 | 0.00001843593794715056 | 0.00001843593794713455 |
| 30 | $0.5846426652345042361310^{-7}$ | $0.5846426638915782682410^{-7}$ |
| 40 | $0.1854025800003477154910^{-9}$ | $0.1854020557057607958110^{-9}$ |
| 50 | $0.5879509107770131156610^{-12}$ | $0.5869540320787337587810^{-12}$ |
| 60 | $0.1922782986051643082010^{-14}$ | $0.6426989414949361226410^{-15}$ |

TABLE II: THE $p(i, j)^{\prime} s$ FOR $\rho=0.2$

| $j \backslash i$ | 0 | $1 \quad 2$ | 2 | 3 |  | 4 | 5 | 6 | 7 | 8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | $8.1710^{-1}$ |  |  |  |  |  |  |  |  |  |
| 1 | $8.1710^{-2}$ | $1.510^{-2}$ |  |  |  |  |  |  |  |  |
| 2 | $7.1210^{-4}$ | $8.510^{-4} \quad 1.6$ | $1.610^{-4}$ |  |  |  |  |  |  |  |
| 3 | $3.5210^{-6}$ | $4.210^{-6} \quad 8.5$ | $8.510^{-6}$ | 1.61 |  |  |  |  |  |  |
| 4 | $1.6810^{-8}$ | $2.010^{-8} \quad 4.0$ | $4.010^{-8}$ | 8.51 |  | $1.610^{-8}$ |  |  |  |  |
| 5 | $7.9910^{-11}$ | $9.510^{-11} \quad 1.9$ | $1.910^{-10}$ | 4.01 | -10 | $8.510^{-10}$ | $1.610^{-10}$ |  |  |  |
| 6 | $3.81{ }^{10^{-13}}$ | $4.510^{-13} \quad 9.2$ | $9.210^{-13}$ | 1.91 | ${ }^{-12}$ | $4.010^{-12}$ | $8.510^{-12}$ | $10^{-12}$ |  |  |
| 7 | $1.8210^{-15}$ | $2.110^{-15} 4.3$ | $4.310^{-15}$ | 9.21 | -15 | $1.910^{-14}$ | $4.010^{-14}$ | $810^{-14}$ | $10^{-14}$ |  |
| 8 | $2.0810^{-17}$ | $1.210^{-17} \quad 2.1$ | $2.110^{-16}$ | 4.31 | -17 | $9.010^{-17}$ | $1.910^{-16}$ | $410^{-16}$ | $810^{-16}$ | $10^{-16}$ |
| 9 | $1.2410^{-17}$ | $2.410^{-18}$ | $4.510^{-19}$ | 8.01 | -20 | $1.310^{-20}$ | $1.910^{-21}$ | $210^{-22}$ | $210^{-23}$ |  |
| 10 | $1.2410^{-17}$ | $2.410^{-18} \quad 4.3$ | $4.310^{-19}$ | 7.21 | -20 | $1.010^{-20}$ | $1.310^{-21}$ | $110^{-22}$ |  |  |
| 11 | $1.2510^{-17}$ | $2.310^{-18} 4.0$ | $4.010^{-19}$ | 6.11 | -20 | $7.810^{-21}$ | $7.110^{-22}$ |  |  |  |
| 12 | $1.2710^{-17}$ | $2.310^{-18} \quad 3.4$ | $3.410^{-19}$ | 4.61 | -20 | $4.210^{-21}$ |  |  |  |  |
| 13 | $1.2910^{-17}$ | $2.110^{-18} \quad 2.8$ | $2.810^{-19}$ | 2.61 | -20 |  |  |  |  |  |
| 14 | $1.3310^{-17}$ | $1.810^{-18} \quad 1$ | $1.710^{-19}$ |  |  |  |  |  |  |  |
|  | $\epsilon$ | $\epsilon_{0}$ | $\epsilon_{1}$ |  | $N$ | the compu | d Sum |  |  |  |
|  | $10^{-12}$ | $0.98110^{-12}$ | 2 0.109 | $0^{-20}$ | 20 | 1.00000000 | 000 |  |  |  |

TABLE III: THE $p(i, j)^{\prime} s$ FOR $\rho=1.9$

| $i \backslash j$ | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0 | 1.99 | $10^{-2}$ |  |  |  |  |  |  |  |  |
| 1 |  |  |  |  |  |  |  |  |  |  |
| 1 | 1.89 | $10^{-2}$ | 2.85 | $10^{-2}$ |  |  |  |  |  |  |
| 2 |  |  |  |  |  |  |  |  |  |  |


| $\epsilon$ | $\epsilon_{0}$ | $\epsilon_{1}$ | $N$ | the computed Sum |
| :--- | :--- | :--- | :--- | :--- |
| $10^{-5}$ | $0.210^{-6}$ | $1.5 \quad 10^{-5}$ | 226 | 1.00000 |

## V. Conclusion

This new numerical method is very simple for use and easily implemented for computation. The tools used are classical linear algebra and simplex method for which the software is available for solving big linear systems of inequalities and a great number of variables. Just a formal calculus (in terms of $\left.\left\{y_{i}^{(N)} ; 1 \leq i \leq\left[\frac{N}{2}\right]\right\}\right)$ of $\left\{r_{N}(i, j) ; i+j=N\right\}$ and the use of the simplex algorithm ( known as one of the most efficient algorithm regarding it's complexity) for solving the system of inequalities $\left\{r_{N}(i, j) \geq 0 ; i+j=N\right\}$. This will permit us to get the unknown steady state probabilities with a desired precision. For small values of $\rho$ this method seems faster compared with the other ones. More $\rho$ is close 0 more the number of operations is greater (see for example [13]). An other point of interest is the adaptation of the method for other cases (due to the same structure of the steady state balance equations) of the shortest queue problem.

## References

[1] Adan, I.J.B.F. and Wessels, J. and Zijm W.H.M. , Analysis of the symmetric shortest queueing problem, Stoch. Models, 1990 6, 691-713
[2] Cohen, J.W. and Boxma, O.J. , Boundary Problems in Queuing Systems Analysis, North- Holland, Amsterdam, 1983
[3] Flatto, L. and McKean, H.P. , Two queues in parallel, Comm. Pure. Appl. Math, 1977 30, 255-263
[4] Gertsbakh, I. , The shorter queue problem: A numerical study using the matrix geometric solution, European J. Operation Research, 1984 15, 374-381
[5] Haight, F.A. , Two queues in parallel, Biometrika, 1958 48, 401-410
[6] Halfin, S. , The shortest queue problem, J. Appl. Probab., 1985 22, 865878.
[7] Kingman, J.F.C. . Two similar queues in parallel, Ann. Math. Stat., 1961 32, 1314-1323.
[8] Lardjane, T. and Messaci, R. , On a new numerical computation of the steady state solution of two infinite server parallel queues, Applied Mathematical Sciences, 2011 Vol 5, 78, 3875-3891.
[9] Neuts, M. F. , Matrix geometric Solutions in Stochastic Models. John Hopkins University Press, Baltimore, 1980
10] Tarabia, A.M.K., Analysis of two queues in parallel with jockeying and restricted capacities, Appl. Math.Modell., 2008 32(5), 802-810
[11] Wang, P. and Locker, V.F. , Steady state distributions of parallel queues INFOR, 2001 39(1).
[12] Yao, H. and Knessl, C. , On the infinite server shortest queue problem: symmetric case, Stochastic Models, 2005 21(1), 101-132
[13] Zhao, Y. and Grassman, W.K. , A numerically stable algorithm for two server queue model, Queueing Systems, 1991, 8, 59-79


[^0]:    T. Lardjane and R. Messaci are with the Faculty of Mathematics and Laboratory MSTD, U. S. T. H. B , B. P. 32. El Alia Bab Ezzouar Algiers, Algeria.
    e-mail: tlardjane@hotmail.com, lrabmes@yahoo.fr

